# Quasi-Poisson regression models for radiation dose estimation from biomarkers 

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## Radiation biodosimetry

- Radiation incident leading to (potentially) exposed individuals
- Contracted radiation dose can be estimated retrospectively by exploiting the radiation-induced change in certain biomarkers
- 'Gold-standard': Dicentric chromosomes (resulting from unsuccessful DNA-damage response)
- However: time- and work-intensive and expensive methodology



## $\gamma-\mathrm{H} 2 \mathrm{AX}$ as radiation biomarker

- Biomarkers based on proteins have recently emerged as a quicker and cheaper alternative
- The H2AX-histone responds to radiation-induced double strand breaks with phosphorylation, in this state then referred to as $\gamma-$ H2AX
- $\gamma-\mathrm{H} 2 \mathrm{AX}$ foci can be counted manually (immunofluorescence microscopy) or automated (flow cytometers)



## H2AX calibration data from PHE

Foci 'yield' (that is foci/cell, out of 500 sample cells) versus design dose:


- Strong (linear?) dose-response relationship; strong decay from 1h to 24 h after exposure
- H2AX-based dose estimation has to happen within 24 hours of exposure!
- Considerable variation, so Uncertainty Quantification crucial


## Calibration curve estimation

- For fixed time after exposure, calibration data $\left(x_{i}, y_{i j}\right), j=1, \ldots, n_{i}$, with $y_{i j}$ "yield of the $j$-th sample (of $n$ cells) for dose $x_{i}$ ".
- Count data (Poisson) regression with 'identity link'

$$
\mu_{i} \equiv E\left(y_{i j} \mid x_{i}\right)=A+B x_{i}
$$

(Quadratic models also considered, but discarded)

- Huge overdispersion ( $\hat{\phi} \approx 60$ ! $)$



## Quasi-Poisson regression

- Overdispersed Poisson model for $Y_{i j}=n y_{i j}$,

$$
E\left(Y_{i j} \mid x_{i}\right)=n A+B\left(n x_{i}\right) ; \quad \operatorname{Var}\left(Y_{i j} \mid x_{i}\right)=\phi E\left(Y_{i j} \mid x_{i}\right)
$$

Score equations for this model (with $\mu_{i}=A+B x_{i}$ ),

$$
\frac{1}{\phi} \sum_{i=1}^{d} \sum_{j=1}^{n_{i}}\binom{1}{x_{i}}\left(Y_{i j}-n \mu_{i}\right) / \mu_{i}=\binom{0}{0}
$$

... so the estimates of $A$ and $B$ do not depend on $\phi$ !

- However, standard errors do depend on $\phi$, namely

$$
S E(\hat{A})=\sqrt{\hat{\phi}} \mathrm{SE}_{P}(\hat{A}) ; \quad \mathrm{SE}(\hat{B})=\sqrt{\hat{\phi}} \mathrm{SE}_{P}(\hat{B})
$$

## Estimating dispersion

- 'Quick-and-dirty', from estimated Poisson model

$$
\hat{\phi}=\frac{\text { Deviance }}{N-2}
$$

- We use

$$
\hat{\phi}=\frac{1}{N-2} \sum_{i=1}^{d} \sum_{j=1}^{n_{i}} \frac{\left(Y_{i j}-n \hat{\mu}_{i}\right)^{2}}{n \hat{\mu}_{i}}=\frac{n}{N-2} \sum_{i=1}^{d} \sum_{j=1}^{n_{i}} \frac{\left(y_{i j}-\hat{\mu}_{i}\right)^{2}}{\hat{\mu}_{i}}
$$

where $\hat{\mu}_{i}=\hat{A}+\hat{B} x_{i}$ and $N=\sum_{i=1}^{d} n_{i}$.

- Each additional covariate will increase the value ' 2 ' in the denominator by 1 .


## Estimating dispersion: Warning!

- Dispersion estimates can become poor if data are aggregated.

Dispersion estimates before and after aggregation over $\mathbf{x}$

- Simulation example: 10000 data sets of size 100 from $y_{i} \sim \operatorname{NB1}\left(\mu_{i}, 2\right)$, with
$\mu_{i}=2+5 x_{i}$ and
$x_{i} \sim \operatorname{Bin}(2,0.5)$.
- That is, $\phi=3$ known.
- Even though the average of estimates $\hat{\phi}$ are almost identical (2.96 and 2.98, resp.), each individual dispersion estimate from the aggregated data is close to useless.


## Dispersion from calibration data

- Our calibration data are 'half-aggregated' (over cells; not over dose).
- Estimated dispersions:

| $\hat{\phi}$ | 1 h | 24 h |
| :--- | :---: | :---: |
| 'quick-and-dirty' | 57.91 | 54.26 |
| GLM-based | 59.56 | 57.55 |



- $\hat{\phi} \approx 60$ appears a reasonable assumption.


## Dose estimation

- Inverse regression: For a new yield $y_{*}$, one has

$$
x_{*}=\frac{y_{*}-\hat{A}}{\hat{B}}
$$



- UQ via delta-method:

$$
\begin{aligned}
S E^{2}\left(x_{*}\right) & =\left(\frac{\partial x_{*}}{\partial \hat{A}}\right)^{2} S E^{2}(\hat{A})+\left(\frac{\partial x_{*}}{\partial \hat{B}}\right)^{2} S E^{2}(\hat{B})+\left(\frac{\partial x_{*}}{\partial y_{*}}\right)^{2} S E^{2}\left(y_{*}\right) \\
& =\frac{1}{\hat{B}^{2}} S E^{2}(\hat{A})+\frac{\left(y_{*}-\hat{A}\right)^{2}}{\hat{B}^{4}} S E^{2}(\hat{B})+\frac{1}{\hat{B}^{2}} \frac{\hat{\phi} y_{*}}{n_{*}}
\end{aligned}
$$

- This accounts for intra- and inter-individual variation, but still requires calibration curve to be 'correct'


## Curve validation

Complication: The calibration curve may vary with laboratory, scorer, equipment etc. Hence, a given calibration curve needs to be validated before use.
Before examining a patient sample, lab should irradiate two reference samples at 0 Gy and 1.5 Gy and compare yields with prediction interval:

- If inside, validated
- If outside, a new calibration curve can be computed from the reference samples which still allows dose estimation, albeit at a higher variance


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## Web applet



## Web applet

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## DoseEstimateH2AX



Main Plot Information


We believe to have solved the questions regarding

- ... the incorporation of overdispersion
- ... the validation of the calibration curve
- ... the quantification of uncertainty in this process

Open questions are

- Is the dispersion a 'universal property of foci counts', or a feature of the scoring mechanism?
- Partial body exposure cannot be easily identified since overdispersion will be present either way


## Use for other (count data) radiation biomarkers

- Micronuclei: Overdispersion usually present but small ( $\approx 1.5$ ), risk of incorrect estimation of $\phi$ possibly larger than benefits. If full count data distributions available then NB model preferable.
- Dicentric chromosomes
- under partial body exposure: Overdispersion is due to zeroinflation; so direct use of ZIP models preferable.
- under densely ionizing radiation: Overdispersion present but situation similar to micronuclei.
- Other protein-based biomarkers, such 53PB1
- here probably useful, but yet to be tried...


## Future directions

- Combinations of Biomarkers

Idea: Use quick and cheap (but potentially high variance) biomarkers such as H2AX for the triage step, and a more precise biomarker (such as the dicentric assay) as a second step, depending on the first outcome. Bayesian approach favorable here.

- Why always estimate 'dose'?

It seems to be an irrevocable standard to always estimate dose. Why not triage directly based on the H2AX count? This removes need for inverse regression, and reduces uncertainties. Could be dealt with easily through ordinal logistic regression.

Ainsbury EA et al (2017). Uncertainty of fast biological radiation dose assessment for emergency response scenarios. International Journal of Radiation Biology 93, 127-135.

Einbeck J et al (2017). On the Use of Random Effect Models for Radiation Biodosimetry. In: Extended Abstracts Fall 2015. Ainsbury EA et al, Research Perspectives CRM Barcelona 7, 89-94, Springer.

Einbeck J et al (2018). A statistical framework for radiation dose estimation from the $\gamma-\mathrm{H} 2 \mathrm{AX}$ assay. PLoS ONE 13(11):e0207464.

Oliveira M et al (2016). Zero-inflated regression models for radiationinduced chromosome aberration data: A comparative study. Biometrical Journal 58, 259-79.

