## Multilevel Monte Carlo for elliptic SPDEs

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## Outline

- standard Monte Carlo simulation
- multilevel Monte Carlo simulation
- elliptic SPDE application
- conclusions

In many applications want to estimate  $\mathbb{E}[P(\omega)]$  where  $\omega \in \Omega$  is an infinite-dimensional random variable.

- computational finance:
  - $\omega$  represents  $W_t$  the driving Brownian motion in an SDE (stochastic differential equation)
  - *P* is the financial payoff function
- simulation of oil reservoirs & nuclear waste repositories:
  ω represents k(x), the diffusivity in an elliptic SPDE

$$-\nabla \cdot \left(k(x)\,\nabla p\right) = 0$$

*P* might be the flux of oil or contaminants across some boundary

In MC simulation we estimate the expectation using

$$\widehat{Y} = N^{-1} \sum_{n=1}^{N} \widehat{P}(\omega^{(n)})$$

where  $\omega^{(n)}$  are N independent samples

Note there are two sources of error here:

- *sampling error* due to the finite number of samples
- **•** bias because  $\widehat{P}(\omega)$  is an approximation to  $P(\omega)$  due to
  - discretisation error (finite timesteps, finite grid size)
  - finite dimensional approximation to  $\omega$

The mean square error is

$$\begin{split} \mathbb{E}\left[\left(\widehat{Y} - \mathbb{E}[P]\right)^2\right] &= \mathbb{E}\left[\left(\widehat{Y} - \mathbb{E}[\widehat{Y}] + \mathbb{E}[\widehat{Y}] - \mathbb{E}[P]\right)^2\right] \\ &= \mathbb{E}\left[\left(\widehat{Y} - \mathbb{E}[\widehat{Y}]\right)^2\right] + \left(\mathbb{E}[\widehat{Y}] - \mathbb{E}[P]\right)^2 \\ &= \mathbb{V}[\widehat{Y}] + \left(\mathbb{E}[\widehat{Y}] - \mathbb{E}[P]\right)^2 \\ &= N^{-1}\mathbb{V}[\widehat{P}] + \left(\mathbb{E}[\widehat{P}] - \mathbb{E}[P]\right)^2 \end{split}$$

- first term is due to sampling error
- second term is due to bias

To achieve RMS accuracy of  $\varepsilon$  requires:

$$N = O(\varepsilon^{-2})$$

• bias 
$$= O(\varepsilon)$$

In a *d*-dimensional SPDE application with grid spacing *h*, if the bias is  $O(h^{\alpha})$  then need  $h = O(\varepsilon^{1/\alpha})$ , and total cost is  $O(\varepsilon^{-(2+d/\alpha)})$ , assuming efficient multigrid solution

(very challenging because of very rough coefficients – Graham & Scheichl)

To get acceptable accuracy in 3D applications may need 10,000 simulations on a  $128^3$  grid  $\implies$  very expensive

The multilevel objective is to greatly reduce this cost:

	$\alpha = 1$		$\alpha = 2$	
dim	MC	MLMC	MC	MLMC
1	$\varepsilon^{-3}$	$\varepsilon^{-2}$	$\varepsilon^{-2.5}$	$\varepsilon^{-2}$
2	$\varepsilon^{-4}$	$\varepsilon^{-2}(\log \varepsilon)^2$	$\varepsilon^{-3}$	$\varepsilon^{-2}(\log \varepsilon)^2$
3	$\varepsilon^{-5}$	$\varepsilon^{-3}$	$\varepsilon^{-3.5}$	$\varepsilon^{-2.5}$

How? Use multigrid philosophy:

- fine grid accuracy at coarse grid cost
- geometric sequence of grids

*but* no iteration in Monte Carlo simulation?

Consider Monte Carlo simulations with different levels of refinement, l = 0, 1, ..., L, with level L being the finest.

If  $\widehat{P}_l$  is the approximation of *P* on level *l*, then

$$\mathbb{E}[\widehat{P}_L] = \mathbb{E}[\widehat{P}_0] + \sum_{l=1}^L \mathbb{E}[\widehat{P}_l - \widehat{P}_{l-1}].$$

Idea is to independently estimate each of the terms on the r.h.s., in a way which minimises the overall variance for a fixed computational cost.

Finest level is still the same, but will use very few samples at that level.

Simplest estimator for  $\mathbb{E}[\widehat{P}_l - \widehat{P}_{l-1}]$  for l > 0 is

$$\widehat{Y}_{l} = N_{l}^{-1} \sum_{n=1}^{N_{l}} \left( \widehat{P}_{l}^{(n)} - \widehat{P}_{l-1}^{(n)} \right)$$

using same stochastic sample  $\omega^{(n)}$  for both levels

Variance is  $N_l^{-1}V_l$  where  $V_l = \mathbb{V}[\widehat{P}_l - \widehat{P}_{l-1}]$ 

Key point:  $V_l$  gets progressively smaller as l increases because  $\hat{P}_l, \hat{P}_{l-1}$  both accurately approximate P for same  $\omega$ 

If  $C_l$  is cost of one sample on level l, the variance of the combined estimator is  $\sum_{l=0}^{L} N_l^{-1} V_l$  and its computational cost is  $\sum_{l=0}^{L} N_l C_l$  so the variance is minimised for fixed cost by choosing  $N_l \propto \sqrt{V_l/C_l}$ , and then the cost on level l is proportional to  $N_l C_l \propto \sqrt{V_l C_l}$ 

To make RMS error  $\varepsilon$ 

shoose constant of proportionality so variance is  $\frac{1}{2} \varepsilon^2$ 

• choose 
$$L$$
 so that  $\left(\mathbb{E}[\widehat{P}_l] - \mathbb{E}[P]\right)^2 < \frac{1}{2}\varepsilon^2$ 

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### **MLMC Theorem**

If there exist independent estimators  $\widehat{Y}_l$  based on  $N_l$  Monte Carlo samples, each costing  $C_l$ , and positive constants  $\alpha, \beta, \gamma, c_1, c_2, c_3$  such that  $\alpha \geq \frac{1}{2} \min(\beta, \gamma)$  and

i)  $\left| \mathbb{E}[\hat{P}_l - P] \right| \leq c_1 2^{-\alpha l}$ ii)  $\mathbb{E}[\hat{Y}_l] = \begin{cases} \mathbb{E}[\hat{P}_0], & l = 0\\ \mathbb{E}[\hat{P}_l - \hat{P}_{l-1}], & l > 0 \end{cases}$ iii)  $\mathbb{V}[\hat{Y}_l] \leq c_2 N_l^{-1} 2^{-\beta l}$ iv)  $C_l \leq c_3 2^{\gamma l}$ 

### **MLMC Theorem**

then there exists a positive constant  $c_4$  such that for any  $\varepsilon < 1$  there exist L and  $N_l$  for which the multilevel estimator

$$\widehat{Y} = \sum_{l=0}^{L} \widehat{Y}_l,$$

has a mean-square-error with bound  $\mathbb{E}\left[\left(\widehat{Y} - E[P]\right)^2\right] < \varepsilon^2$ 

with a computational cost C with bound

$$C \leq \begin{cases} c_4 \varepsilon^{-2}, & \beta > \gamma, \\ c_4 \varepsilon^{-2} (\log \varepsilon)^2, & \beta = \gamma, \\ c_4 \varepsilon^{-2 - (\gamma - \beta)/\alpha}, & 0 < \beta < \gamma. \end{cases}$$

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# **Papers**

- My first paper (Operations Research, 2006 2008) applied idea to SDE path simulation, and proved slightly less general form of the theorem
- Second paper (MCQMC 2006) improved multilevel variance convergence using better discretisation
- Third paper with D. Higham & X. Mao (*Finance and Stochastics, 2009*) performed numerical analysis of discretisation in first paper
- New paper with K. Debrabant and A. Rößler analyses discretisation in second paper

Multilevel method is a generalisation of two-level control variate method of Kebaier (2005), and related to multilevel parametric integration by Heinrich (2001).

## **Elliptic SPDE**

We consider the elliptic PDE

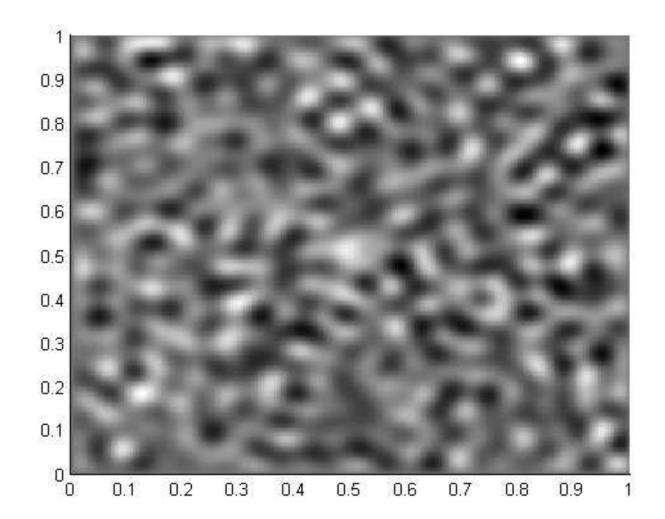
$$-\nabla (k(x,\omega)\nabla p(x,\omega)) = f(x,\omega), \qquad x \in D,$$

with random coefficient  $k(x, \omega)$  and random data  $f(x, \omega)$ .

We model k as a **lognormal random field**, i.e.  $\log k$  is a Gaussian field with mean 0 and covariance function

$$R(x, y) = \sigma^2 \exp\left(-\|x - y\|/\lambda\right)$$

### **Elliptic SPDE**



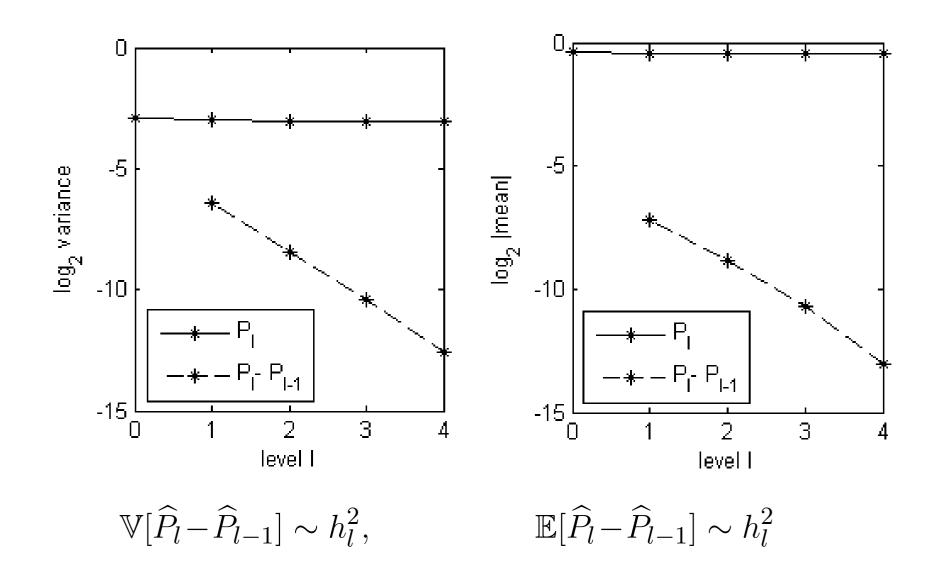
A typical realisation for  $D = [0, 1]^2, \lambda = 0.001, \sigma^2 = 1$ .

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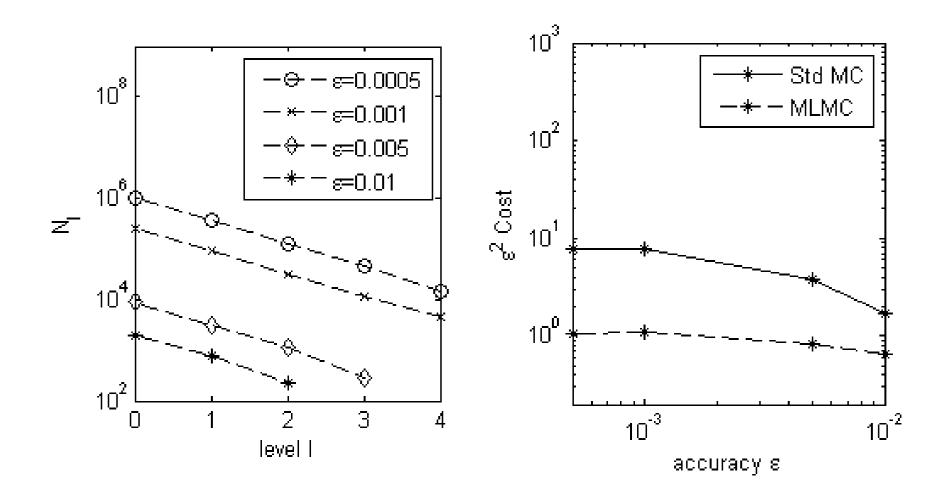
# **Elliptic SPDE**

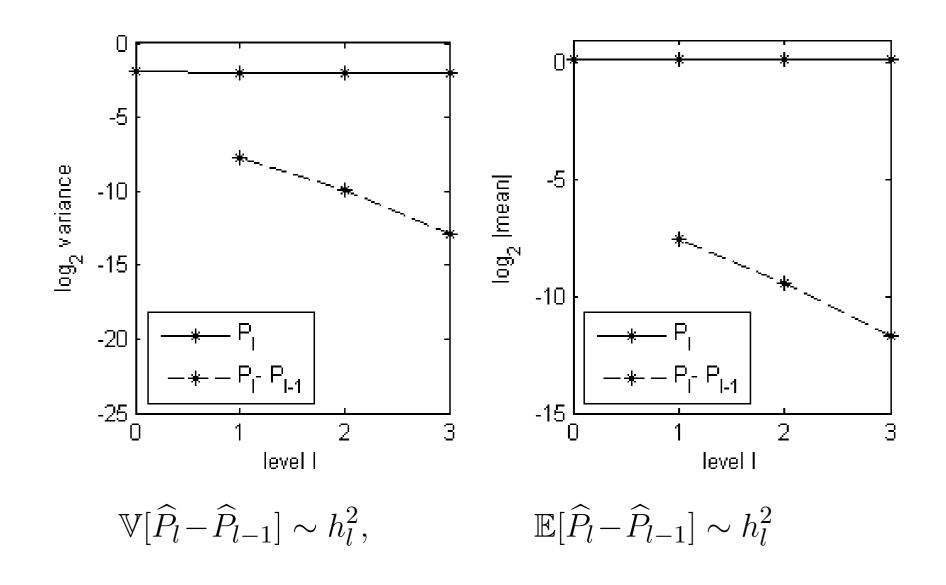
Discretisation:

- cell-centred Finite Volume discretisation on a uniform grid  $\mathcal{T}_h$  for rough coefficients we need to make h very small
- sampling of the random coefficient currently based on truncated Karhunen-Lõeve expansion, evaluated at cell centres – but the method of sampling is not essential to the algorithm
- each level of refinement has twice as many grid points in each direction

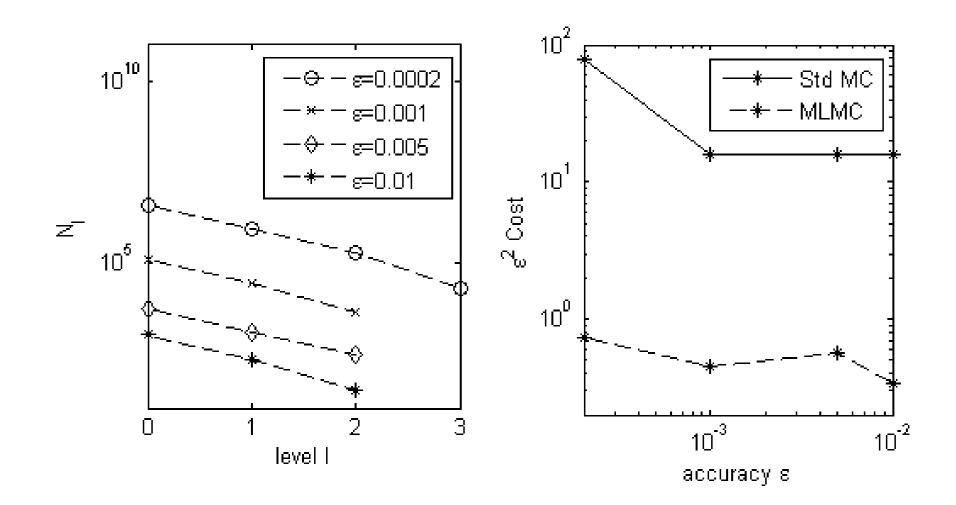


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## Conclusions

- standard Monte Carlo is prohibitively expensive for 2D and 3D elliptic SPDE applications
- multilevel Monte Carlo greatly reduces the cost, making this feasible for engineering applications
- we believe it is a viable competitor to polynomial chaos approach, particularly for applications with minimal spatial correlation
- numerical analysis is very hard, but we're making some headway with finite element analysis, at least to gain insight into its effectiveness
- future work will look at combining the multilevel approach with quasi-Monte Carlo sampling – has been very effective for SDE applications in finance