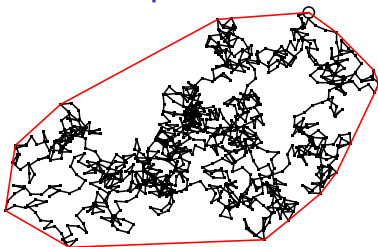


Convex hulls of planar random walks



Andrew Wade

Department of Mathematical Sciences

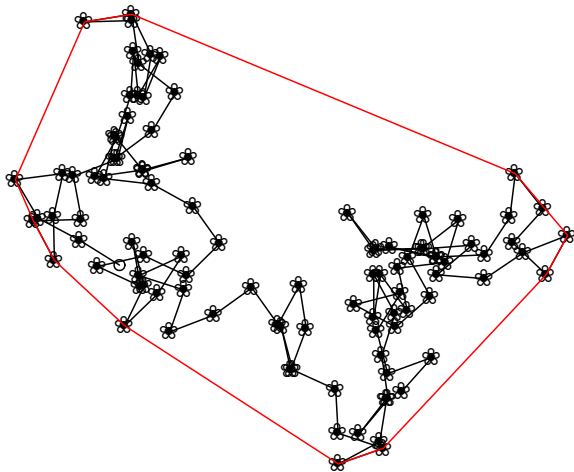


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Joint work with Chang Xu, University of Strathclyde

Introduction

On each of n unsteady steps, a drunken gardener drops a seed. Once the flowers have bloomed, what is the minimum length of fencing needed to enclose the garden? What is its area?



Introduction

Let Z_1, Z_2, \dots be independent, identically distributed random vectors in \mathbb{R}^2 .

The Z_k will be the increments of the **planar random walk** S_n , $n \geq 0$, started at the origin in \mathbb{R}^2 , defined by

$$S_0 = 0, \quad \text{and} \quad S_n = \sum_{k=1}^n Z_k \quad \text{for } n \geq 1.$$

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In particular, the $n \rightarrow \infty$ limit behaviour of the random variables

- L_n = the perimeter length of $\text{hull}(S_0, \dots, S_n)$;
- A_n = the area of $\text{hull}(S_0, \dots, S_n)$.

Introduction

Standing assumption: $\mathbb{E}(\|Z_1\|^2) < \infty$ (sometimes more).

For the **mean drift** vector of the walk we write $\mu = \mathbb{E} Z_1$.

There is going to be a clear distinction between the **zero drift** case ($\mu = 0$) and the **non-zero drift** case ($\|\mu\| > 0$).

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For example, if

$$r_n := \inf\{\|x\| : x \in \mathbb{R}^2 \setminus \text{hull}(S_0, \dots, S_n)\},$$

then, under mild conditions:

- the zero-drift walk visits all angles at arbitrary distances, so $r_n \rightarrow \infty$, i.e., the convex hull tends to the whole of \mathbb{R}^2 ;
- the walk with non-zero drift is transient in a limiting direction, so $\lim_{n \rightarrow \infty} r_n < \infty$.

Introduction

Some heuristics allow us to guess the typical scalings with n of our quantities of interest:

	$\mu = 0$	$\mu \neq 0$
L_n	$n^{1/2}$	n
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We then seek **distributional limit theorems**. That is, for quantity Q_n is there a distributional limit for

$$n^{-\alpha} Q_n \text{ or } n^{-\beta} (Q_n - \mathbb{E} Q_n) ?$$

Given some **self-averaging**, we might seek **laws of large numbers**:

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It turns out that each of the four cases in the table do satisfy a distributional limit theorem: **one** limit distribution is **Gaussian**; the other three are not.

Outline

- 1 Introduction
- 2 Cauchy's formula
- 3 Scaling limits and Brownian hulls
- 4 Scaling limit in the non-zero drift case
- 5 Perimeter length in the non-zero drift case
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Some history

Spitzer & Widom (1961) and **Baxter** (1961) showed that

$$\mathbb{E} L_n = 2 \sum_{k=1}^n \frac{1}{k} \mathbb{E} \|S_k\|.$$

So, under mild conditions:

- the zero-drift case has $\mathbb{E} L_n \asymp \sqrt{n}$;
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- the case with drift has $\mathbb{E} L_n \asymp n$.

Snyder & Steele (1993) showed that

$$\frac{1}{n} \text{Var}(L_n) \leq \frac{\pi^2}{2} \left(\mathbb{E} \|Z_1\|^2 - \|\mu\|^2 \right). \quad (1)$$

Snyder & Steele deduced from (1) the **strong law**

$$\lim_{n \rightarrow \infty} n^{-1} L_n = 2\|\mu\|, \text{ a.s.}$$

Some questions

The work of Snyder & Steele raised some natural questions.

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- Is there a distributional limit theorem for L_n ?
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First, we introduce an important tool from convex geometry:

Cauchy's formula.

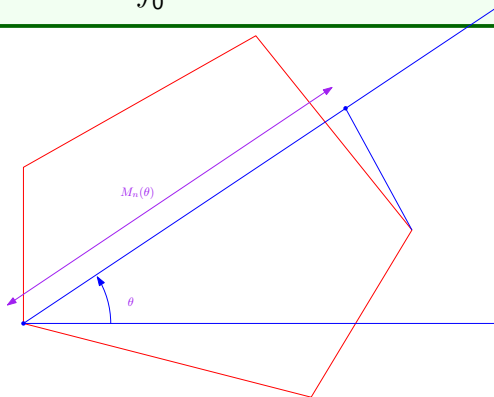
Cauchy's formula

Let $\mathbf{e}_\theta = (\cos \theta, \sin \theta)$, unit vector in direction θ . Set

$$M_n(\theta) = \max_{0 \leq k \leq n} (\mathbf{S}_k \cdot \mathbf{e}_\theta), \quad m_n(\theta) = \min_{0 \leq k \leq n} (\mathbf{S}_k \cdot \mathbf{e}_\theta).$$

Cauchy's perimeter formula from convex geometry:

$$L_n = \int_0^\pi (M_n(\theta) - m_n(\theta)) d\theta.$$



Cauchy's formula

$$L_n = \int_0^\pi (M_n(\theta) - m_n(\theta)) d\theta.$$

A first consequence: classical fluctuation theory for random walk on \mathbb{R} gives

$$\mathbb{E} M_n(\theta) = \sum_{k=1}^n k^{-1} \mathbb{E} [(S_k \cdot \mathbf{e}_\theta)^+],$$

a formula attributed variously to **Kac**, **Hunt**, **Dyson**, and **Chung**, and which can be proved combinatorially, or analytically as a consequence of the **Spitzer–Baxter–Pollaczek** fluctuation theory identities. Then

$$\mathbb{E} L_n = \sum_{k=1}^n k^{-1} \mathbb{E} \int_0^{2\pi} |S_k \cdot \mathbf{e}_\theta| d\theta = 2 \sum_{k=1}^n k^{-1} \mathbb{E} \|S_k\|,$$

which is the **Spitzer–Widom** formula.

Outline

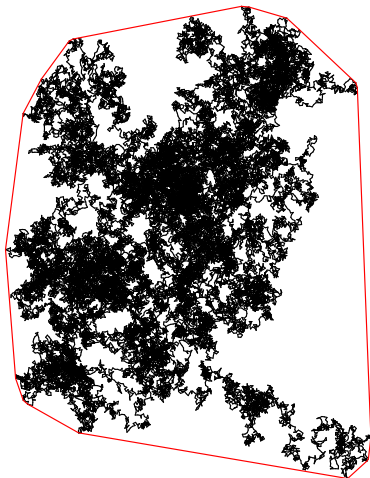
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The zero drift case

Suppose $\mu = 0$. The random walk has **Brownian motion** as its scaling limit.

So one would expect that the convex hull of the random walk is described in the limit by the **convex hull of Brownian motion**. The latter was studied by **Lévy**; more recently by **Ei Bachir** (1983) and others.

We need to know a little about convex hulls of continuous paths, and need to set things up on the right space(s).



Paths and hulls

Consider continuous $f : [0, T] \rightarrow \mathbb{R}^d$ with $f(0) = 0$; say $f \in C_d^0$.
(T is not very important—enough to take $T \equiv 1$.)

With the supremum norm $\rho_\infty(f, g) = \sup_x \|f(x) - g(x)\|$ we get
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The **path** segment (\equiv **interval image**) $f[0, t] = \{f(s) : s \in [0, t]\}$
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\implies hull $f[0, t]$ is compact (by a theorem of **Carathéodory**).

That is, hull $f[0, t]$ is an element of the metric space $(\mathcal{K}_d^0, \rho_H)$ of
compact convex subsets of \mathbb{R}^d containing 0, with the Hausdorff
metric.

Paths and hulls

Metric space $(\mathcal{K}_d^0, \rho_H)$ of compact convex subsets of \mathbb{R}^d containing 0, with the **Hausdorff metric**.

Given $K \in \mathcal{K}_d^0$ and $r > 0$, let $K^r := \{x \in \mathbb{R}^d : \rho(x, K) \leq r\}$.

For $A, B \in \mathcal{K}_d^0$,

$$\rho_H(A, B) \leq r \iff A \subseteq B^r \text{ and } B \subseteq A^r.$$

Lemma 1

For each t , the map $f \mapsto \text{hull } f[0, t]$ is a continuous function from (C_d^0, ρ_∞) to $(\mathcal{K}_d^0, \rho_H)$.

Scaling limit

Given random walk $S_n = \sum_{k=1}^n Z_k$ on \mathbb{R}^d , define

$$X_n(t) := n^{-1/2} (S_{\lfloor nt \rfloor} + (nt - \lfloor nt \rfloor)Z_{\lfloor nt \rfloor + 1}).$$

So for each n , $X_n \in C_d^0$. Let b_t , $t \geq 0$ denote standard Brownian motion on \mathbb{R}^d .

Donsker's Theorem

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$, $\mu = 0$, and $\mathbb{E}(Z_1 Z_1^\top) = I$. Then $X_n \Rightarrow b$ in the sense of weak convergence on (C_d^0, ρ_∞) .

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Note $\text{hull } X_n[0, 1] = n^{-1/2} \text{hull}(S_0, \dots, S_n)$. Then with Lemma 1 and the continuous mapping theorem, we get:

Theorem 2

Under the same conditions,

$$n^{-1/2} \text{hull}(S_0, \dots, S_n) \Rightarrow \text{hull } b[0, 1]$$

in the sense of weak convergence on $(\mathcal{K}_d^0, \rho_H)$.

Functionals

Now take $d = 2$.

The neatest way to define **area** \mathcal{A} and **perimeter length** \mathcal{L} of a set $K \in \mathcal{K}_2^0$ is:

$$\mathcal{A}(K) := |K|, \text{ and } \mathcal{L}(K) := \lim_{r \downarrow 0} \left(\frac{|K^r| - |K|}{r} \right),$$

where $|\cdot|$ is Lebesgue measure; the limit exists by the **Steiner formula** of integral geometry.

In particular,

$$\mathcal{L}(K) = \begin{cases} \mathcal{H}_1(\partial K) & \text{if } \text{int}(K) \neq \emptyset \\ 2\mathcal{H}_1(\partial K) & \text{if } \text{int}(K) = \emptyset \end{cases}$$

where \mathcal{H}_1 is one-dimensional Hausdorff measure.

Functionals

Lemma 3

The maps $K \mapsto \mathcal{A}(K)$ and $K \mapsto \mathcal{L}(K)$ are continuous functions from $(\mathcal{K}_2^0, \rho_H)$ to (\mathbb{R}_+, ρ) .

Note $\mathcal{L}(\text{hull } X_n[0, 1]) = \mathcal{L}(n^{-1/2} \text{hull}(S_0, \dots, S_n)) = n^{-1/2} L_n$;
 $\mathcal{A}(\text{hull } X_n[0, 1]) = \mathcal{A}(n^{-1/2} \text{hull}(S_0, \dots, S_n)) = n^{-1} A_n$.

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Theorem 4

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$, $\mu = 0$, and $\mathbb{E}(Z_1 Z_1^\top) = I$. Then

$$n^{-1/2} L_n \xrightarrow{d} \ell_1, \quad \text{and} \quad n^{-1} A_n \xrightarrow{d} a_1,$$

where $\ell_1 = \mathcal{L}(\text{hull } b[0, 1])$ and $a_1 = \mathcal{A}(\text{hull } b[0, 1])$ are the perimeter length and area, respectively, of the convex hull of planar Brownian motion run for unit time.

Functionals

Some comments.

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- Under mild conditions, we also have that

$$n^{-1} \text{Var} L_n \rightarrow \text{Var} \ell_1, \quad \text{and} \quad n^{-2} \text{Var} A_n \rightarrow \text{Var} a_1.$$

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- We'd like an exact formula for these variances.
Goldman (1996) manages to do an explicit computation for the planar **Brownian bridge**, but it is tricky.
Using Cauchy's formula and a formula of **Rogers & Shepp** (2006) on the correlation of the maxima of correlated one-dimensional Brownian motions, one can show

$$\mathbb{E}[\ell_1^2] = 4\pi \int_{-\pi/2}^{\pi/2} d\theta \int_0^{\infty} du \cos \theta \frac{\cosh(u\theta)}{\sinh(u\pi/2)} \tanh\left(\frac{(2\theta + \pi)u}{4}\right),$$

which gives $\text{Var } \ell_1 \approx 1.0350$.

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Scaling limit in the case with drift

To get a non-degenerate scaling limit, we now must scale space by factor $1/n$ in the **direction of the drift** and by factor $1/\sqrt{n}$ in the **orthogonal direction**.

$$\text{Specifically, let } \psi_n^\mu(x) = \left(\frac{x \cdot \hat{\mu}}{n\|\mu\|}, \frac{x \cdot \hat{\mu}_\perp}{\sqrt{n\sigma_{\mu_\perp}^2}} \right);$$

$$\text{Here } \sigma_{\mu_\perp}^2 = \mathbb{E}[(Z \cdot \hat{\mu}_\perp)^2].$$

Let \tilde{b} denote the process on \mathbb{R}^2 given by $\tilde{b}(t) = (t, w(t))$, where w is standard Brownian motion on \mathbb{R} .

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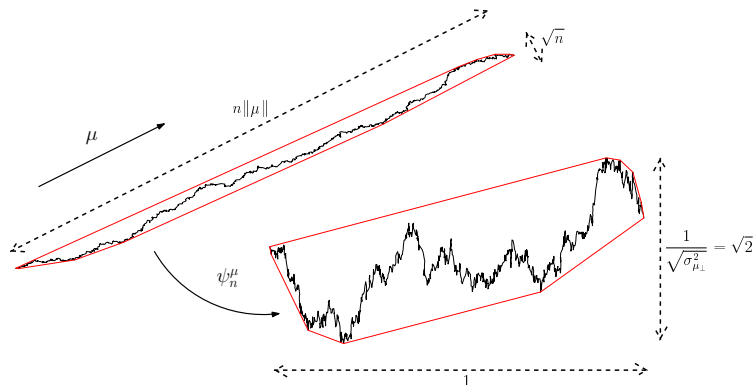
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The analogue of Donsker's theorem is as follows.

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$, $\mu \neq 0$, and $\sigma_{\mu_\perp}^2 > 0$. Then $\psi_n^\mu(X_n) \Rightarrow \tilde{b}$ weakly on $(\mathcal{C}_2^0, \rho_\infty)$.

Scaling limit in the case with drift



The **affine map** ψ_n^{μ} preserves the convex hull, so:

Theorem 5

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$, $\mu \neq 0$, and $\sigma_{\mu_{\perp}}^2 > 0$. Then

$\psi_n^{\mu}(\text{hull}(S_0, \dots, S_n)) \Rightarrow \text{hull } \tilde{b}[0, 1]$ weakly on $(\mathcal{K}_2^0, \rho_H)$.

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Now note $\mathcal{A}(\psi_n^\mu(\text{hull}(S_0, \dots, S_n))) = n^{-3/2} \|\mu\|^{-1} (\sigma_{\mu^\perp}^2)^{-1/2} A_n$.

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Theorem 6

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$$n^{-3/2} \|\mu\|^{-1} (\sigma_{\mu^\perp}^2)^{-1/2} A_n \xrightarrow{d} \tilde{a}_1,$$

where $\tilde{a}_1 = \mathcal{A}(\text{hull } \tilde{b}[0, 1])$.

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Some comments.

- The variable \tilde{a}_1 is **positive** and **non-degenerate**, and so **non-Gaussian**.
- We can show that $\mathbb{E} \tilde{a}_1 = \frac{1}{3}\sqrt{2\pi}$, using an analogue of the Spitzer–Widom formula for **areas** due to **Barndorff-Nielsen & Baxter** (1963).
- Under mild conditions, we also have that

$$n^{-3}\text{Var } A_n \rightarrow \|\mu\|^2 \sigma_{\mu^\perp}^2 \text{Var } \tilde{a}_1.$$

- We can show $\text{Var } \tilde{a}_1 > 0$.
- This scaling limit strategy is no use for the **perimeter length** L_n in the case $\mu \neq 0$, because ψ_n^μ does not act in a sensible way on lengths. . .

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- 1 Introduction
- 2 Cauchy's formula
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- 5 Perimeter length in the non-zero drift case**
- 6 Concluding remarks

Perimeter length in the non-zero drift case

Theorem 7

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$ and $\mu \neq 0$. Then

$$\lim_{n \rightarrow \infty} \frac{1}{n} \text{Var}(L_n) = 4\sigma_\mu^2, \quad \text{where } \sigma_\mu^2 := \mathbb{E}[\left((Z_1 - \mu) \cdot \hat{\mu}\right)^2].$$

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Theorem 8

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$, $\mu \neq 0$, and $\sigma_\mu^2 > 0$. Then

$$\frac{L_n - \mathbb{E} L_n}{\sqrt{\text{Var} L_n}} \xrightarrow{d} \mathcal{N}(0, 1), \quad \text{and} \quad \frac{L_n - \mathbb{E} L_n}{\sqrt{4n\sigma_\mu^2}} \xrightarrow{d} \mathcal{N}(0, 1).$$

Remarks

(i) A little algebra shows $4\sigma_\mu^2 \leq 4(\mathbb{E} \|Z_1\|^2 - \|\mu\|^2)$.

Compare to the **Snyder–Steele** upper bound

$$n^{-1} \text{Var}(L_n) \leq \frac{\pi^2}{2} (\mathbb{E} \|Z_1\|^2 - \|\mu\|^2).$$

I.e., the constant in the Snyder–Steele upper bound is not sharp ($4 < \pi^2/2$).

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I.e., the constant in the Snyder–Steele upper bound is not sharp ($4 < \pi^2/2$).

(ii) $\sigma_\mu^2 = 0$ if and only if $Z_1 - \mu$ is a.s. orthogonal to μ .

This is the case, for instance, if Z_1 takes values $(1, 1)$ or $(1, -1)$, each with probability $1/2$.

In this case Theorem 7 says that $\text{Var}(L_n) = o(n)$.

The Snyder–Steele bound says only that $\text{Var}(L_n) \leq \pi^2 n/2$.

Simulations suggest that actually $\text{Var}(L_n) = O(\log n)$.

Degenerate example

Z_1 takes values $(1, 1)$ or $(1, -1)$, each with probability $1/2$.

This 2-dimensional walk can be viewed as a space-time diagram of a **1-dimensional simple symmetric random walk**:



Interesting combinatorics here, related to the **Bohnenblust–Spitzer** algorithm; see **Steele** (2002).

Behaviour of L_n for this case is an open problem.

Proof idea: Martingale differences

Let $\mathcal{F}_n = \sigma(Z_1, \dots, Z_n)$. Define $D_{n,i} = \mathbb{E}[L_n - L_n^{(i)} \mid \mathcal{F}_i]$.

Lemma 9

$$(i) L_n - \mathbb{E} L_n = \sum_{i=1}^n D_{n,i}.$$

$$(ii) \text{Var}(L_n) = \sum_{i=1}^n \mathbb{E}(D_{n,i}^2).$$

Sketch proof.

As Z_i' is independent of \mathcal{F}_i , $\mathbb{E}[L_n^{(i)} \mid \mathcal{F}_i] = \mathbb{E}[L_n^{(i)} \mid \mathcal{F}_{i-1}] = \mathbb{E}[L_n \mid \mathcal{F}_{i-1}]$.

So $D_{n,i} = \mathbb{E}[L_n \mid \mathcal{F}_i] - \mathbb{E}[L_n \mid \mathcal{F}_{i-1}]$; a standard construction of a martingale difference sequence.

$$\sum_{i=1}^n D_{n,i} = \mathbb{E}[L_n \mid \mathcal{F}_n] - \mathbb{E}[L_n \mid \mathcal{F}_0] = L_n - \mathbb{E} L_n.$$

Now use orthogonality of martingale differences. □

Aside: Upper bounds

Lemma 9 with the conditional Jensen inequality gives:

$$\text{Var}(L_n) \leq \sum_{i=1}^n \mathbb{E}[(L_n - L_n^{(i)})^2].$$

A related result, a version due to **Steele** of the **Efron–Stein inequality**, says

$$\text{Var}(L_n) \leq \frac{1}{2} \sum_{i=1}^n \mathbb{E}[(L_n - L_n^{(i)})^2].$$

It is this latter result that Snyder & Steele used to obtain their upper bound.

Cauchy formula revisited

We need to study $D_{n,i} = \mathbb{E}[L_n - L_n^{(i)} \mid \mathcal{F}_i]$.

We have the Cauchy formula for L_n , and similarly for $L_n^{(i)}$, so that

$$L_n - L_n^{(i)} = \int_0^\pi \Delta_{n,i}(\theta) d\theta,$$

where

$$\Delta_{n,i}(\theta) = \left(M_n(\theta) - M_n^{(i)}(\theta) \right) - \left(m_n(\theta) - m_n^{(i)}(\theta) \right),$$

where, similarly to before,

$$M_n^{(i)}(\theta) = \max_{0 \leq k \leq n} (\mathcal{S}_k^{(i)} \cdot \mathbf{e}_\theta), \quad m_n^{(i)}(\theta) = \min_{0 \leq k \leq n} (\mathcal{S}_k^{(i)} \cdot \mathbf{e}_\theta).$$

Proof idea: Control of extrema

We want to understand the relationship between $M_n(\theta)$, $m_n(\theta)$ and $M_n^{(i)}(\theta)$, $m_n^{(i)}(\theta)$ (resampled versions).

WLOG suppose $\mathbb{E} Z_1 = \mu \mathbf{e}_{\pi/2} = (0, \mu)$, where $\mu > 0$.

Then for each fixed θ , $S_n \cdot \mathbf{e}_\theta$ is a **one-dimensional** random walk.

Indeed, $S_n \cdot \mathbf{e}_\theta = \sum_{k=1}^n Z_k \cdot \mathbf{e}_\theta$, with **mean increment** $\mathbb{E}[Z_1 \cdot \mathbf{e}_\theta] = \mathbb{E}[Z_1] \cdot \mathbf{e}_\theta = \mu \sin \theta$, which is **positive** for $\theta \in (0, \pi)$.

So, with high probability, the max $M_n(\theta)$ will be achieved nearby step n while the min $m_n(\theta)$ will be achieved nearby step 0.

To formalize this needs the **strong law of large numbers**, plus some care (need some **uniformity** in θ).

Proof idea: Control of extrema

Lemma 10

With high probability, $\Delta_{n,i}(\theta) \approx (Z_i - Z'_i) \cdot \mathbf{e}_\theta$ for (almost) all i .

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So $m_n(\theta) = m_n^{(i)}(\theta)$, and

$$M_n^{(i)}(\theta) = M_n(\theta) + (Z'_i - Z_i) \cdot \mathbf{e}_\theta.$$

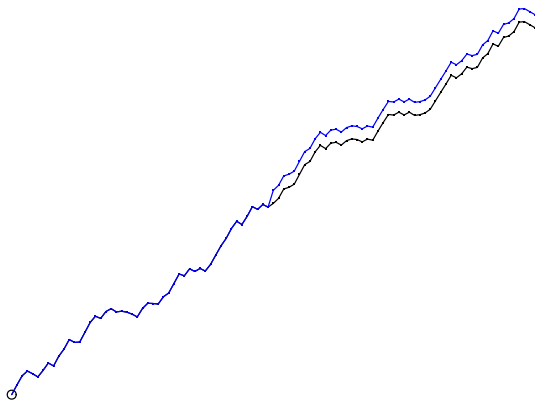
See the picture!

Proof idea: Control of extrema

So $m_n(\theta) = m_n^{(i)}(\theta)$, and

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See the picture!



Finishing the proofs

Up to technical details, we sketched the fact that

$$D_{n,i} = \mathbb{E} [L_n - L_n^{(i)} \mid \mathcal{F}_i] \approx \int_0^\pi \mathbb{E} [(Z_i - Z'_i) \cdot \mathbf{e}_\theta \mid \mathcal{F}_i] d\theta.$$

Here Z_i is \mathcal{F}_i -measurable and Z'_i is independent of \mathcal{F}_i , so

$$\mathbb{E} [(Z_i - Z'_i) \cdot \mathbf{e}_\theta \mid \mathcal{F}_i] = (Z_i - \mu) \cdot \mathbf{e}_\theta.$$

Doing the integral gives

$$D_{n,i} = \mathbb{E} [L_n - L_n^{(i)} \mid \mathcal{F}_i] \approx 2(Z_i - \mu) \cdot \hat{\mu}.$$

Finishing the proofs

Formalizing the analysis we get:

Theorem 11

Suppose $\mathbb{E}(\|Z_1\|^2) < \infty$ and $\mu \neq 0$. Then

$$n^{-1/2} \left| L_n - \mathbb{E} L_n - 2 \sum_{i=1}^n (Z_i - \mu) \cdot \hat{\mu} \right| \rightarrow 0, \text{ in } L^2.$$

So, perhaps surprisingly, $L_n - \mathbb{E} L_n$ is well-approximated by a sum of **i.i.d.** random variables.

Theorems 7 and 8 now follow from Theorem 11 easily enough.

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Concluding remarks

The assumption that the Z_i are identically distributed is not essential to the main argument.

For example, let $G_n = \frac{1}{n+1} \sum_{i=0}^n S_i = \sum_{i=1}^n \frac{n+1-i}{n+1} Z_i$.

G_0, G_1, \dots is the **centre-of-mass** process associated with S_0, S_1, \dots

By convexity, $\text{hull}(G_0, \dots, G_n) \subseteq \text{hull}(S_0, \dots, S_n)$.

If L_n^G is the perimeter length of $\text{hull}(G_0, \dots, G_n)$, then the statement of Theorem 11 applies to L_n^G in place of L_n with $\frac{n+1-i}{n+1} Z_i$ in place of Z_i .

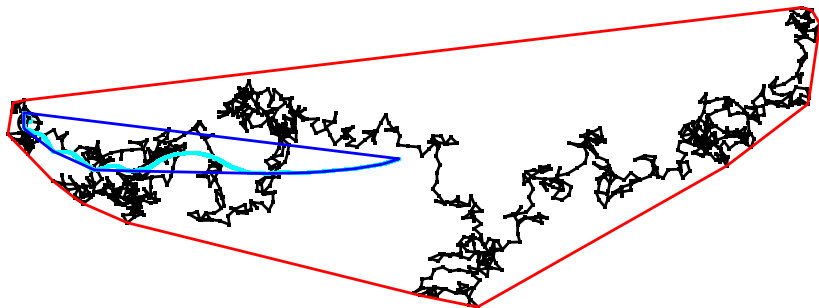
In particular, the analogue of Theorem 7 says that

$$\lim_{n \rightarrow \infty} \frac{1}{n} \text{Var}(L_n^G) = 4\sigma_\mu^2/3,$$

where σ_μ^2 is the same as before.

Concluding remarks

A picture:



References

- O. BARNDORFF-NIELSEN & G. BAXTER, Combinatorial lemmas in higher dimensions, *Trans. Amer. Math. Soc.* **108** (1963) 313–325.
- G. BAXTER, A combinatorial lemma for complex numbers, *Ann. Math. Statist.* **32** (1961) 901–904.
- M. EL BACHIR, *L'enveloppe convexe du mouvement brownien*, Ph.D. thesis, Université Toulouse III, 1983.
- A. GOLDMAN, Le spectre de certaines mosaïques poissoniennes du plan et l'enveloppe convexe du pont brownien, *Probab. Theory Relat. Fields* **105** (1996) 57–83.
- S.N. MAJUMDAR, A. COMTET, & J. RANDON-FURLING, Random convex hulls and extreme value statistics, *J. Stat. Phys.* **138** (2010) 955–1009.
- L.C.G. ROGERS & L. SHEPP, The correlation of the maxima of correlated Brownian motions, *J. Appl. Probab.* **43** (2006) 880–883.
- T.L. SNYDER & J.M. STEELE, Convex hulls of random walks, *Proc. Amer. Math. Soc.* **117** (1993) 1165–1173.
- F. SPITZER & H. WIDOM, The circumference of a convex polygon, *Proc. Amer. Math. Soc.* **12** (1961) 506–509.
- J.M. STEELE, The Bohnenblust–Spitzer algorithm and its applications, *J. Comput. Appl. Math.* **142** (2002) 235–249.
- A.R. WADE & C. XU, Convex hulls of planar random walks with drift, *Proc. Amer. Math. Soc.* **143** (2015) 433–445.
- A.R. WADE & C. XU, Convex hulls of random walks and their scaling limits, *Stoch. Proc. Appl.* **125** (2015) 4300–4320.