

Confidence Bands for Smoothers

Ben Powell

Durham

May 12, 2010

- 1 Notation
- 2 Variance Estimates
- 3 Confidence Bands
- 4 Confidence Regions in Smoother Space
- 5 Model Selection

Notation

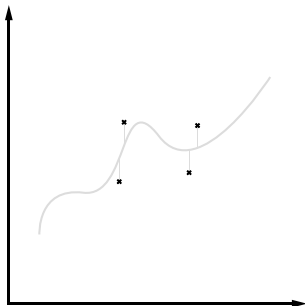
- $\mathbf{y} = \mathbf{f} + \epsilon$
- $\hat{\mathbf{f}} = \mathbf{S}\mathbf{y}$

Where can we use variance estimates?

- Confidence intervals
- Prediction intervals
- Curiosity

Local estimates

$$\hat{\sigma}_L^2 = \frac{1}{2(n-1)} \sum_{i=2}^n (y_i - y_{i-1})^2$$



Local estimates

$\hat{\sigma}_L^2$ is positively biased

$$\begin{aligned} E[\hat{\sigma}_L^2] &= E\left[\frac{1}{2(n-1)} \sum_{i=2}^n (y_i - y_{i-1})^2\right] \\ &= \frac{1}{2(n-1)} \sum_{i=2}^n E[(y_i - y_{i-1})^2] \\ &= \frac{1}{2(n-1)} \sum_{i=2}^n E[(f_i + \epsilon_i)^2 - 2(f_i + \epsilon_i)(f_{i-1} + \epsilon_{i-1}) + \dots] \\ &\vdots \\ &= \sigma^2 + \frac{1}{2(n-1)} \sum_{i=2}^n (f_i - f_{i-1})^2 \end{aligned}$$

Local estimates

(At least) two ways around this:

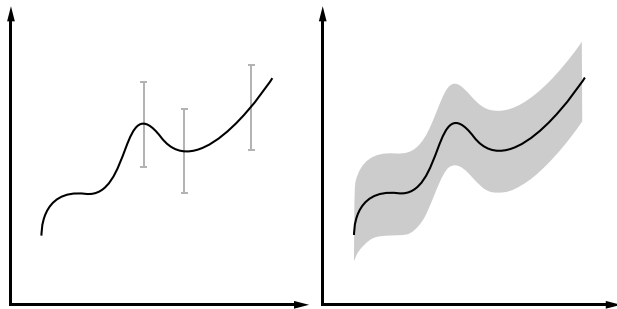
- Introduce different lags.
- Try: $\hat{\sigma}_R^2 = \frac{1}{n-2} \sum_{i=2}^{n-1} (y_i - \tilde{y}_i)^2$ where \tilde{y}_i is the fitted value at x_i for a linear regression on the three closest data points.

Global estimates

$$\hat{\sigma}_\lambda^2 = \frac{1}{n - p} \sum_{i=1}^n (y_i - \hat{f}_i)^2,$$

p = degrees of freedom for smoother

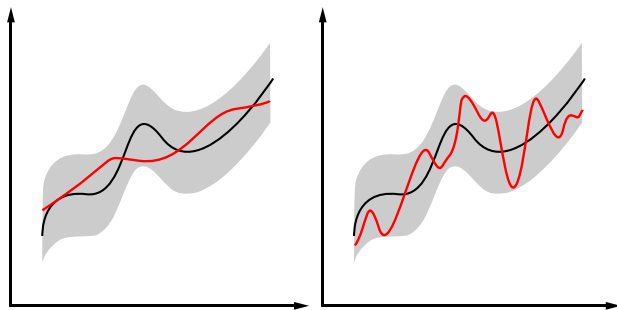
Confidence bands



(a) Conf. intervals

(b) Conf. band

Confidence bands



(c) sample smoother 1

(d) sample smoother 2

Notation

$$\hat{\mathbf{f}} = \mathbf{S}\mathbf{y}$$

$$\mathbf{y} = \mathbf{f} + \boldsymbol{\epsilon}$$

$$\boldsymbol{\epsilon} \sim N(\mathbf{0}, \mathbf{I}_n\sigma^2)$$

$$\epsilon^t \epsilon \sigma^{-2} \sim \chi_n^2$$

$$\epsilon^t \mathbf{S}^t (\mathbf{S} \mathbf{S}^t \sigma^2)^{-1} \mathbf{S} \epsilon \sim \chi_{\text{rank}(\mathbf{S})}^2$$

$$(\mathbf{S} \epsilon)^t (\mathbf{S} \mathbf{S}^t \sigma^2)^{-1} (\mathbf{S} \epsilon) \sim \chi_n^2$$

$$(\mathbf{S} f + \mathbf{S} \epsilon - \mathbf{S} f)^t (\mathbf{S} \mathbf{S}^t \sigma^2)^{-1} (\mathbf{S} f + \mathbf{S} \epsilon - \mathbf{S} f) \sim \chi_n^2$$

$$\left(\hat{\mathbf{f}} - \mathbf{S} f \right)^t (\mathbf{S} \mathbf{S}^t \sigma^2)^{-1} \left(\hat{\mathbf{f}} - \mathbf{S} f \right) \sim \chi_n^2$$

$$\hat{\epsilon} = \mathbf{y} - \hat{\mathbf{f}} = (\mathbf{I} - \mathbf{S})\mathbf{y} = (\mathbf{I} - \mathbf{S})\mathbf{f} + (\mathbf{I} - \mathbf{S})\epsilon$$

$$\mathbf{S}\mathbf{f} = \mathbf{f} ?$$

$$\hat{\epsilon} = (\mathbf{I} - \mathbf{S})\epsilon$$

Detour

Recall...

$$\begin{aligned} F(\mathbf{g}) &= (\mathbf{y} - \mathbf{f})^t (\mathbf{y} - \mathbf{f}) + \lambda \mathbf{f}^t \mathbf{K} \mathbf{f} \\ &= (\mathbf{y} - \phi \beta)^t (\mathbf{y} - \phi \beta) + \lambda (\phi \beta)^t \mathbf{K} \phi \beta \\ &\dots \\ &\Rightarrow \hat{\beta} = (\phi^t \phi + \lambda \phi^t \mathbf{K} \phi)^{-1} \phi^t \mathbf{y} \end{aligned}$$

- Using Sherman-Morrison-Woodbury formula and ϕ symmetric

$$\begin{aligned} \mathbf{S} &= \phi (\phi^t \phi + \lambda \phi^t \mathbf{K} \phi)^{-1} \phi^t \\ &= (\mathbf{I} - \lambda \mathbf{K})^{-1} \end{aligned}$$

$$\begin{aligned}\mathbf{Sf} &= \mathbf{f} \\ \Rightarrow (\mathbf{I} - \lambda\mathbf{K})\mathbf{f} &= \mathbf{f} \\ \Rightarrow \mathbf{Kf} &= \mathbf{0}\end{aligned}$$

...back to business

$$\hat{\epsilon}^t \hat{\epsilon} = \epsilon^t (\mathbf{I} - \mathbf{S})^t (\mathbf{I} - \mathbf{S}) \epsilon$$

$$\frac{\hat{\epsilon}^t \hat{\epsilon}}{\sigma^2} = \frac{\epsilon^t (\mathbf{I} - \mathbf{S})^t (\mathbf{I} - \mathbf{S}) \epsilon}{\sigma^2} = ?$$

$$\chi_{tr(\mathbf{I} - \mathbf{S} - \mathbf{S}^t + \mathbf{S}^t \mathbf{S})}^2$$

$$n - p = \text{tr}(\mathbf{I} - \mathbf{S} - \mathbf{S}^t + \mathbf{S}^t \mathbf{S})$$

$$\hat{\sigma}_\lambda^2 = \frac{\hat{\epsilon}^t \hat{\epsilon}}{n - p} \sim \frac{\sigma^2 \chi_{n-p}^2}{n - p}$$

Constructing the test

$$\begin{aligned} & (\hat{\mathbf{f}} - \mathbf{S}f)^t (\mathbf{S}\mathbf{S}^t \hat{\sigma}_\lambda^2)^{-1} (\hat{\mathbf{f}} - \mathbf{S}f) \\ & \sim \frac{n-p}{\sigma^2 \chi_{n-p}^2} (\hat{\mathbf{f}} - \mathbf{S}f)^t (\mathbf{S}\mathbf{S}^t)^{-1} (\hat{\mathbf{f}} - \mathbf{S}f) \\ & \sim \frac{n-p}{\chi_{n-p}^2} \chi_n^2 \\ & \sim \frac{n-p}{\chi_{n-p}^2} (\chi_{n-p}^2 + \chi_p^2) \\ & \sim (n-p) \left(1 + \frac{\chi_p^2}{\chi_{n-p}^2}\right) \\ & \sim (n-p) + pF_{p,n-p} \qquad \text{(test distribution)} \end{aligned}$$

Constructing the test

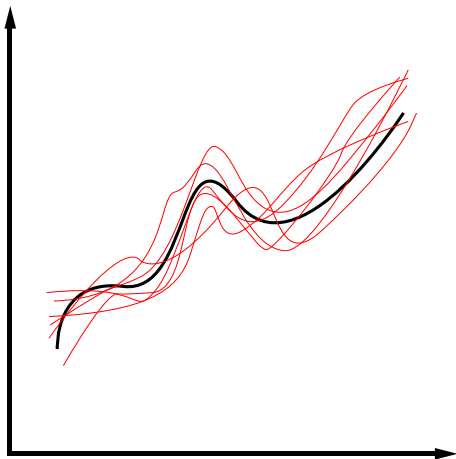
Nice result but assumptions may cause test to be misleading

- $\mathbf{Sf} = \mathbf{f}$
- $(\mathbf{I} - \mathbf{S})$ behaves like $(\mathbf{I} - \mathbf{H})$ from linear regression.

Viewing functions from the interval

- Draw MVN vectors from $N(\hat{\mathbf{f}}, \hat{\sigma}_\lambda^2)$
- Accept or reject depending on the test distribution statistic.

Viewing functions from the interval



Model Selection

Use F-test completely analogously to the linear regression (ANOVA) context.

If the fitted model $\hat{\mathbf{f}}_2 = \mathbf{S}_2\mathbf{y}$ at least contains all the terms of $\hat{\mathbf{f}}_1 = \mathbf{S}_1\mathbf{y}$, with residual sums of squares RSS_2 and RSS_1 and effective degrees of freedom p_2 and p_1 , then:

$$\frac{(RSS_1 - RSS_2)/(p_2 - p_1)}{RSS_2/(n - p_2)} \sim F_{p_2 - p_1, n - p_2}$$

References

- Hastie, T.J. & Tibshirani, R.J. (1990), **Generalized Additive Models**, Chapman and Hall.
- Green, P.J. & Silverman, B.W. (1994), **Nonparametric regression and generalized linear models**, Chapman and Hall.
- Rice, J. (1984), *Bandwidth Choice for Nonparametric Regression*, *Annals of Statistics* **12(4)**, pp 1215-1230.
- Tong, T. & Wang, Y. (2005), *Estimating residual variance in nonparametric regression using least squares*, *Biometrika* **92(4)**, pp 821-830.