

### Hints to exercises 11.1, 11.2, 11.3, 11.4 (week 17)

- 11.1 At time  $y$ , the price of the security equals  $S(y)$ . Suppose you own an option to buy one share of the stock at a price  $Ke^{uy}$  at time  $y$ . The option expires after an additional time  $t - y$ .

If you exercise the option at  $y < t$  you realize the amount  $S(y) - Ke^{uy}$ , which is worth  $(S(y) - Ke^{uy})e^{(t-y)r}$  at time  $t$ . Now, instead of exercising the option early, suppose you sell the stock short and purchase the stock at time  $t$ , either by paying the market price at that time or by exercising your option and paying  $Ke^{ut}$ , whichever is less expensive. Hence, in this case, your payoff is worth at time  $t$   $S(y)e^{(t-y)r} - \min(Ke^{ut}, S(t))$ .

As  $S(y)e^{(t-y)r} - \min(Ke^{ut}, S(t)) \geq (S(y) - Ke^{uy})e^{(t-y)r}$  (\*) you never exercise the call option earlier if  $u \geq r$ . To see why (\*) is true, first suppose  $\min(Ke^{ut}, S(t)) = Ke^{ut}$ . As  $uy + tr - yr \geq ut$  (i.e.  $(t - y)(r - u) \geq 0$ ), we have that  $Ke^{ut} \leq Ke^{uy+tr-yr}$  and consequently  $S(y)e^{(t-y)r} - Ke^{ut} \geq (S(y) - Ke^{uy})e^{(t-y)r}$  as required.

Now if  $\min(Ke^{ut}, S(t)) = S(t)$  then  $S(t) < Ke^{ut} \leq Ke^{uy+tr-yr}$  and hence  $S(y)e^{(t-y)r} - S(t) > S(y)e^{(t-y)r} - Ke^{uy+tr-yr} = (S(y) - Ke^{uy})e^{(t-y)r}$  as required.

- 11.2 Section 11.5 explains how to simulate a sequence of end-of-day prices and from these the payoff is easily calculated. Section 11.6.1 discusses the use of antithetic variables to reduce variation and improve efficiency.

- 11.3  $V = (S_d(n) - K)^+$  so  $E(V) = E((S_d(n) - K)^+) = e^{rn/N} C(S_d(0), n/N, K, \sigma, r)$   
 where  $C = S_d(0)\Phi(\omega) - Ke^{-rn/N}\Phi(\omega - \sigma\sqrt{n/N})$   
 and  $\omega = rn/N + \frac{1}{2}\sigma^2\frac{n}{N} - \log(K/S_d(0))/\sigma\sqrt{n/N}$ .

- 11.4 (a)  $W = Y + \sum_{i=1}^n c_i(X_i - \mu_i)$  so

$$\begin{aligned} \text{Var}(W) &= \text{Var}(Y + \sum_{i=1}^n c_i(X_i - \mu_i)) \\ &= \text{Var}(Y) + \text{Var}(\sum_{i=1}^n c_i(X_i - \mu_i)) + 2\text{Cov}(Y, \sum_{i=1}^n c_i(X_i - \mu_i)) \\ &= \text{Var}(Y) + \sum_{i=1}^n c_i^2 \text{Var}(X_i) + 2 \sum_{i=1}^n c_i \text{Cov}(Y, X_i). \end{aligned}$$

- (b) We must differentiate with respect to the  $c_i$ .

$$\begin{aligned} \frac{\partial \text{Var}(W)}{\partial c_i} &= 2c_i \text{Var}(X_i) + 2\text{Cov}(Y, X_i) = 0, \quad i = 1, \dots, n \\ \Rightarrow c_i &= -\text{Cov}(Y, X_i)/\text{Var}(X_i) \quad i = 1, \dots, n \\ \text{and } \frac{\partial^2 \text{Var}(W)}{\partial c_i^2} &= 2\text{Var}(X_i) \geq 0, \quad \frac{\partial^2 \text{Var}(W)}{\partial c_i \partial c_j} = 0. \end{aligned}$$

Matrix of second derivatives is diagonal with  $2\text{Var}(X_i)$  as diagonal elements and hence is positive definite so this stationary point is a local minimum.